TOWARD A MATHEMATICAL HOLOGRAPHIC PRINCIPLE

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ABSTRACT. In work started in [17] and continued in this paper our objective is to study selectors of multivalued functions which have interesting dynamical properties, such as possessing absolutely continuous invariant measures. We specify the graph of a multivalued function by means of lower and upper boundary maps τ_1 and τ_2 . On these boundary maps we define a position dependent random map $R_p = \{\tau_1, \tau_2; p, 1 - p\}$, which, at each time step, moves the point x to $\tau_1(x)$ with probability p(x) and to $\tau_2(x)$ with probability 1 - p(x). Under general conditions, for each choice of p, R_p possesses an absolutely continuous invariant measure with invariant density f_p . Let τ be a selector which has invariant density function f. One of our objectives is to study conditions under which p(x) exists such that R_p has f as its invariant density function. When this is the case, the long term statistical dynamical behavior of a selector can be represented by the long term statistical behavior of a random map on the boundaries of G. We refer to such a result as a mathematical holographic principle. We present examples and study the relationship between the invariant densities attainable by classes of selectors and the random maps based on the boundaries and show that, under certain conditions, the extreme points of the invariant densities for selectors are achieved by bang-bang random maps, that is, random maps for which $p(x) \in \{0, 1\}$.

1. INTRODUCTION

A function $\tau : X \to X$ maps every $x \in X$ to only one point $y = \tau(x)$. There are applications where this is not the case. For example, in economics, a consumer's action may not manifest itself in a uniquely determined process. This is also common in game theory. To study such applications, we need a new analytical tool, the multivalued function or correspondence as it is called in the mathematical economics literature [4,8]. A multivalued function $\Gamma : X \Rightarrow X$ is a function from Xto the set 2^X of all subsets of X. The graph of Γ is the set: $G = \{(x, y) \in X \times X | y \in \Gamma(x)\}$. Such maps have important applications in rigorous numerics [21], in economics [5, 13, 19], in dynamical systems [2,3,10], chaos synchronization [29], and in differential inclusions [11]. Once Γ is specified one considers maps $\tau : X \to X$ with $\tau(x) \in \Gamma(x)$. Such maps are called selectors. Establishing the existence of continuous selectors in topological spaces has been an area of active interest for more than 60 years [1, 12, 26, 30–32]. In the setting of chaotic dynamical systems, however, selectors possessing measure theoretic, rather than topological, properties are of paramount importance.

In this paper we study selectors of a multivalued function that possess absolutely continuous invariant measures (acims) and relate their dynamics to the dynamics of random maps that are based solely on the boundaries of the graph G defining the multivalued function. We refer to such a property as holographic. Figure 1 shows an example of a region G with boundary maps and a selector τ . Very loosely, the Holographic Principle claims that information in the interior of a black hole can be described by information on its boundary. In this note we attempt to establish a basis for an analogous dynamical system result: let τ be a selector with values in G, and with probability density function (pdf) f. The main problem we address is: under what conditions on the boundary maps $\{\tau_1, \tau_2\}$ and on τ , can we find a probability function p(x) such that the resulting random map $R_p = \{\tau_1, \tau_2; p, 1 - p\}$

Date: February 12, 2014.

²⁰⁰⁰ Mathematics Subject Classification. 37A05, 37H99, 60J05.

Key words and phrases. multivalued functions; selector; absolutely continuous invariant measure; mathematical holographic principle.

The research of the authors was supported by NSERC grants. The research of Z. Li is also supported by NNSF of China (No. 11161020 and No. 11361023).

has f as its invariant pdf? When this is the case, the long term statistical behavior of the map τ inside G is represented by the long term statistical behavior of a random map defined only on the boundaries of G. Such a result qualifies to be referred to as a mathematical Holographic Principle. Gaining insight into this dynamical holographic problem is the main objective of this paper.



FIGURE 1. Lower and upper boundary maps τ_1 , τ_2 and a selector $\boldsymbol{\tau}$.

We shall use random maps [14–16], where the probabilistic weights associated with the maps are functions of position as opposed to being constants in the standard random map framework [29]. We define a position dependent random map as follows: let $R_p = \{\tau_1, \tau_2; p, 1 - p\}$, where τ_1 and τ_2 are maps of the unit interval and p and 1 - p are position dependent probabilities, that is, $0 \le p(x) \le 1$. At each step, the random map moves the point x to $\tau_1(x)$ with probability p(x) and to $\tau_2(x)$ with probability 1 - p(x). Of interest to this paper is the following result [15,16]: for fixed $\{\tau_1, \tau_2\}$, R_p can have different invariant pdf's, depending on the choice of the (weighting) functions p. Let f_k be an invariant density of τ_k , k = 1, 2. It is shown in [15] that for any positive constants c_k , k = 1, 2, there exists a system of weighting probability function p such that the density $f = c_1 f_1 + c_2 f_2$ is invariant under the random map $R_p = \{\tau_1, \tau_2; p, 1 - p\}$, where $p = \frac{c_1 f_1}{c_1 f_1 + c_2 f_2}$. (It is assumed that 0/0 = 0.)

We now list some of the problems studied in this paper by section:

When can a pdf of a selector be realized as a pdf of a random map based on the boundary maps? In Section 2 we attempt to gain insight by considering several examples illustrating that sometimes this is possible but in similar cases it is not. In Section 3, Theorem 1 and Corollary 1 state that any convex combination of pdf's of boundary maps can be realized both as the pdf of a selector and as the pdf of a random map. In Section 4 we use a result from [22] to study functional equations whose solutions guarantee that a given selector with pdf f can be achieved by a random map on the boundaries. In Section 5 we present an example where the boundary maps each have a global attracting invariant measure that is singular. Surprisingly, however, we will show that a selector that has Lebesgue invariant measure can be achieved by a position dependent random map on the boundary maps.

In Section 6, we ask: when can a pdf of a random map based on the boundary maps be realized as a pdf of a selector? Theorem 3 proves that this always holds in the case of piecewise expanding boundary maps which are of the same monotonicity on partition intervals.

Section 7 establishes a continuity theorem that is used in Section 8, where we give a characterization of the extreme points of the set of pdf's of all random maps based on the boundary maps. In the last two sections we present a summary of the examples and concluding remarks.

2. Gaining insight

In this section we present a number of examples, both positive and negative, as we explore the problem: when a pdf of a selector can be realized as a pdf of a random map based on the boundary maps. Example 1 presents a positive general solution for the case when the selector is the triangle map and the boundary maps have two branches, the first quadratic, the second linear.

Example 1. In Figure 2 a simple region G defined by piecewise monotonic maps consisting of 2 pieces each, where on [a, 1] the graphs of both boundary maps are the same and defined by the slope $\frac{1}{1-a}$ map. The selector we consider is the triangle map, whose pdf is f = 1. We will show that f can be achieved by a position dependent random map on the boundary maps. In fact, in this example we can find the exact form of p(x).



FIGURE 2. Maps τ_1 and τ_2 from Example 1.

FIGURE 3. Invariant function p for Example 1.

Let us consider $\tau_i : [0, a] \to [0, 1]$, i = 1, 2, increasing with $\tau_i(0) = 0$, $\tau_i(a) = 1$ satisfying $\tau_1(x) \le x/a$, $\tau_2(x) \ge x/a$ and define the map $\tau_{21} = \tau_2^{-1} \circ \tau_1 : [0, a] \to [0, a]$. τ_{21} is increasing with $\tau_{21}(0) = 0$, $\tau_{21}(a) = a$, and satisfies $\tau_{21}(x) < x$ on the open interval (0, a).

We consider the special case where τ_1 and τ_2 are quadratic:

$$\tau_1(x) = a_1 x^2 + b_1 x$$
, $\tau_2(x) = a_2 x^2 + b_2 x$,

where $a_i = (1 - ab_i)/a^2$, $i = 1, 2, 0 < b_1 < 1/a < b_2 < 2/a$. It follows that

$$\tau_{21}(x) = \frac{1}{2(1-ab_2)} \cdot \left(-a^2b_2 + \sqrt{a^4b_2^2 + 4(1-ab_2)[(1-ab_1)x^2 + a^2b_1x]} \right)$$

The Frobenius-Perron operator (see [10, 15]) associated with the random map $R = \{\tau_1, \tau_2; p, 1-p\}$ is

$$(P_R f)(x) = \frac{p(\tau_1^{-1}(x))f(\tau_1^{-1}(x))}{\tau_1'(\tau_1^{-1}(x))} + \frac{\left(1 - p(\tau_2^{-1}(x))\right)f(\tau_2^{-1}(x))}{\tau_2'(\tau_2^{-1}(x))} + (1 - a)f(1 - (1 - a)x).$$
(1)

If there exists a probability function p such that R preserves the pdf f = 1, i.e. $(P_R 1)(x) = 1$ a.e., then it follows from Equation (1) that

$$p(\tau_1^{-1}(x)) = p(\tau_2^{-1}(x)) \frac{\tau_1'(\tau_1^{-1}(x))}{\tau_2'(\tau_2^{-1}(x))} - \frac{\tau_1'(\tau_1^{-1}(x))}{\tau_2'(\tau_2^{-1}(x))} + a\tau_1'(\tau_1^{-1}(x)).$$
(2)

After introducing the new variable $y = \tau_1^{-1}(x)$ and replacing y with x, equation (2) becomes the functional equation

$$p(x) = p(\tau_{21}(x)) \cdot \tau'_{21}(x) - [\tau'_{21}(x) - a \cdot \tau'_{1}(x)] \quad . \tag{3}$$

Now we need to verify that p satisfies

$$0 \leq p \leq 1$$
 .

The linear function p(x) = A + Bx satisfies (3) if we let

$$A = \frac{b_1(1 - ab_2)}{b_1 - b_2} \quad , \quad B = \frac{2(1 - ab_1)(1 - ab_2)}{(b_1 - b_2)a^2} \quad .$$

Let a = 1/3, $b_1 = 0.5$ and $b_2 = 4.5$. Then A = 0.0625 and B = 1.875. The maps τ_1 , τ_2 and the function p are shown in Figures 2 and 3, respectively. Note that p(x) is defined on [0, a], and is arbitrary on (a, 1] since both maps are the same on (a, 1].

Example 2 presents a positive solution for the specific case of a selector and semi-Markov boundary maps with five branches.

Example 2. Consider boundary maps which are piecewise linear semi-Markov maps as in [16]. Consider τ_1 and τ_2 as defined below and shown in Fig.4. We can show that the selector defined by $\tau(x) = 5x \mod 1$, which has pdf f = 1, can be attained by a position dependent random map.

$$\tau_1(x) = \begin{cases} 2x , & \text{for } 0 \le x \le 0.1 ; \\ 8x - 0.6 , & \text{for } 0.1 \le x < 0.2 ; \\ 5x \mod 1 , & \text{for } 0.2 \le x \le 1 . \end{cases} \quad \tau_2(x) = \begin{cases} 8x , & \text{for } 0 \le x \le 0.1 ; \\ 2x + 0.6 , & \text{for } 0.1 \le x < 0.2 ; \\ 5x \mod 1 , & \text{for } 0.2 \le x \le 1 . \end{cases}$$



FIGURE 4. A selector whose pdf can be achieved by a random map defined by the boundary maps.

Let us consider the random map $R = \{\tau_1, \tau_2; p, 1-p\}$, where

$$p(x) = \begin{cases} 0.2 , & \text{for } 0 \le x < 0.1 ; \\ 0.8 , & \text{for } 0.1 \le x \le 0.2 \end{cases}$$

and arbitrary on the interval (0.2, 1]. Let us use the matrix operator [10] which corresponds to the Frobenius-Perron operator associated with the random map R. Denote the induced matrices of τ_1 and τ_2 by M_1 and M_2 , respectively. We denote the probability function p by the vector $[p_1, p_2, \ldots, p_6]$, where $p_1 = 0.2$, $p_2 = 0.8$, and p_k , $k = 3, \ldots, 6$, are arbitrary. Let q equal $[1 - p_1, 1 - p_2, \ldots, 1 - p_6]$. Now the induced matrix of the map R [16] is

$$M = diag(p)M_1 + diag(q)M_2 .$$

Thus,

$$M = \begin{bmatrix} 1/5 & 1/5 & 1/10 & 1/10 & 1/10 & 0 \\ 0 & 0 & 1/10 & 1/10 & 1/10 & 1/5 \\ 1/5 & 1/5 & 1/5 & 1/5 & 1/5 & 1/5 \\ 1/5 & 1/5 & 1/5 & 1/5 & 1/5 & 1/5 \\ 1/5 & 1/5 & 1/5 & 1/5 & 1/5 & 1/5 \\ 1/5 & 1/5 & 1/5 & 1/5 & 1/5 & 1/5 \end{bmatrix}.$$

It can be checked that the left invariant eigenvector is [1, 1, 1, 1, 1], which implies the random map R preserves Lebesgue measure, the acim of the selector $\tau(x) = 5x \mod 1$.

Example 3 shows that with the same boundary maps as in Example 2 and another selector (which has an invariant pdf), it is impossible.

Example 3. In this example, τ_1 and τ_2 are the same as in Example 2, but the selector τ is defined by:

$$\boldsymbol{\tau}(x) = \begin{cases} 6x , & for \ 0 \le x \le 0.1 ; \\ 4x + 0.2 , & for \ 0.1 \le x < 0.2 ; \\ 5x \bmod 1 , & for \ 0.2 \le x \le 1 . \end{cases}$$



FIGURE 5. A selector whose pdf cannot be achieved by a random map defined by the boundary maps.

The graph is shown in Fig.5. The invariant density of τ is

$$f(x) = \frac{30}{31}\chi_{[0,0.6]}(x) + \frac{65}{62}\chi_{[0.6,1]}(x).$$
(4)

Let

$$p(x) = \begin{cases} p_1(x) , & \text{for } 0 \le x \le 0.2 ; \\ 1 , & \text{for } 0.2 < x \le 1 . \end{cases}$$

Let q(x) = 1 - p(x), for $x \in [0, 1]$. We define a random map $R = \{\tau_1, \tau_2; p(x), q(x)\}$ whose Frobenius-Perron operator [15] is given by

$$(P_R f)(x) = \frac{p_1(\frac{x}{2})f(\frac{x}{2})}{2}\chi_{[0,0,2]}(x) + \frac{p_1(\frac{x}{8} + \frac{3}{40})f(\frac{x}{8} + \frac{3}{40})}{8}\chi_{[0,2,1]}(x)$$

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$$+ \frac{\left(1 - p_1\left(\frac{x}{8}\right)\right)f\left(\frac{x}{8}\right)}{8}\chi_{_{[0,0,8]}}(x) + \frac{\left(1 - p_1\left(\frac{x}{2} - \frac{3}{10}\right)\right)f\left(\frac{x}{2} - \frac{3}{10}\right)}{2}\chi_{_{[0,8,1]}}(x) \\ + \frac{1}{5}\sum_{i=3}^{6}f\left(\tau_{1,i}^{-1}(x)\right).$$

If f(x) defined in (4) is an invariant density of R, then we must have

$$(P_R f)(x) = f(x), \ a.e.$$

Thus, we obtain

$$\begin{aligned} \frac{15}{31}p_1(\frac{x}{2})\chi_{[0,0,2]}(x) &+ \frac{15}{124}p_1(\frac{x}{8} + \frac{3}{40})\chi_{[0,2,1]}(x) \\ &- \frac{15}{124}p_1(\frac{x}{8})\chi_{[0,0,8]}(x) - \frac{15}{31}p_1(\frac{x}{2} - \frac{3}{10})\chi_{[0,8,1]}(x) \\ &= \frac{5}{124}\chi_{[0,0,6]}(x) + \frac{15}{124}\chi_{[0,6,0,8]}(x) - \frac{15}{62}\chi_{[0,8,1]}(x). \end{aligned}$$

Now, if $x \in [0.6, 0.8]$ *, we have*

$$p_1(\frac{x}{8} + \frac{3}{40}) - p_1(\frac{x}{8}) = 1.$$
(5)

Thus, $p_1(z) = 0$ for $z \in [0.075, 0.1]$ while $p_1(z) = 1$ for $z \in [0.15, 0.175]$. Now consider $x \in [0.9, 0.95]$, we have

$$p_1\left(\frac{x}{8} + \frac{3}{40}\right) = 4p_1\left(\frac{x}{2} - \frac{3}{10}\right) - 2.$$
 (6)

Since $\frac{x}{2} - \frac{3}{10} \in [0.15, 0.175]$, it follows from Equation (5) that

$$p_1(\frac{x}{8} + \frac{3}{40}) = 2, \ x \in [0.9, 0.95],$$

which is a contradiction to p_1 being a probability function.

In summary, the ability to represent the long term statistical dynamics of an arbitrary selector by the long term statistical dynamics of a position dependent random map based on the boundaries is a very sensitive matter.

3. Classes of selectors which can be "represented" by random maps on the boundaries

Without loss of generality, we assume that τ_1 and τ_2 have a common partition: $a_0 = 0 < a_1 < a_2 < \cdots < a_m = 1$. Let $I_j = [a_{j-1}, a_j], j = 1, 2, 3, \ldots, m$. On each interval $I_j, \tau_{1,j} := \tau_1|_{I_j}$ and $\tau_{2,j} := \tau_2|_{I_j}$ share the same monotonicity, where we understand $\tau_{1,j}$ and $\tau_{2,j}$ as the natural extensions of pieces of τ_1 and τ_2 , respectively.

Let us assume that τ_1 and τ_2 preserve measures μ_1 and μ_2 , respectively. Let $F^{(1)}, F^{(2)}$ be the distribution functions of measures μ_1, μ_2 , respectively, i.e. $F^{(i)}(x) = \mu_i([0, x]), i = 1, 2$. Let $\mu = \lambda \mu_1 + (1 - \lambda)\mu_2, 0 < \lambda < 1$, and let F be the distribution function of μ , i.e. $F(x) = \mu([0, x])$. Let \mathcal{T} be the class of piecewise C^1 maps from I into I whose graphs belong to G. In [17], the following theorem is proved.

Theorem 1. Let $\tau_1, \tau_2 \in \mathcal{T}$ and assume they possess continuous invariant distribution functions $F^{(1)}$ and $F^{(2)}$. Then, for any convex combination $F = \lambda F^{(1)} + (1 - \lambda)F^{(2)}$, $0 < \lambda < 1$, there exists a piecewise monotonic selector $\boldsymbol{\tau}, \tau_1 \leq \boldsymbol{\tau} \leq \tau_2$, preserving the distribution function F.

Moreover, the formula of each piece of τ is given. For any interval $[a, b] \subseteq [0, 1]$, given a monotone continuous function $h : [a, b] \to [0, 1]$, we define its extended inverse as follows. Let

$$h^{\max} = \max\{h(x)|x \in [a,b]\}, \text{ and } h^{\min} = \min\{h(x)|x \in [a,b]\}.$$

Depending on h is increasing or decreasing, its extended inverse is defined as

$$\overline{h^{-1}}(x) = \begin{cases} a \ , \ \text{for} \ x \in [0, h^{\min}] \ ; \\ h^{-1}(x) \ , \ \text{for} \ x \in [h^{\min}, h^{\max}] \ ; \\ b \ , \ \text{for} \ x \in [h^{\max}, 1] \ ; \end{cases} \qquad \overline{h^{-1}}(x) = \begin{cases} b \ , \ \text{for} \ x \in [0, h^{\min}] \ ; \\ h^{-1}(x) \ , \ \text{for} \ x \in [h^{\min}, h^{\max}] \ ; \\ a \ , \ \text{for} \ x \in [h^{\max}, 1] \ , \end{cases}$$

respectively. We define the extended inverse of each branch of au by

$$\overline{\tau_j^{-1}}(x) = F^{-1}\left(\lambda F^{(1)}(\overline{\tau_{1,j}^{-1}}(x)) + (1-\lambda)F^{(2)}(\overline{\tau_{2,j}^{-1}}(x))\right),\tag{7}$$

where j = 1, 2, 3, ..., m. τ defined in this way, after the vertical segments are removed, has the same number of branches as τ_1 and τ_2 .

In equation (7), if $\lambda = 0$, then our selector τ is τ_2 ; if $\lambda = 1$, then the selector τ is τ_1 . Therefore, as λ varies from 0 to 1, we have a decomposition of G into a pairwise disjoint union of curves, which can be though of as a kind of "foliation" of G.

The above Theorem 1 and the Theorem 4 in [15] together establish the following corollary.

Corollary 1. Let τ be the selector that preserves the density $f = \lambda f_1 + (1 - \lambda) f_2$. Then there exists a probability function p such that random map $R = \{\tau_1, \tau_2; p, 1 - p\}$ also preserves f.

From Theorem 6 in [17] and Corollary 1, it follows that we can consider a more general case:

Theorem 2. Let $\tau_i = h_i^{-1} \circ \Lambda \circ h_i$, i = 1, 2, where Λ is a triangle map, h_1 and h_2 are two diffeomorphisms. Define a diffeomorphism $h = \lambda h_1 + (1 - \lambda)h_2$. Then the map $\tau = h^{-1} \circ \Lambda \circ h$ is a selector between τ_1 and τ_2 , and thus its pdf can be preserved by a random map defined by τ_1 and τ_2 .

4. FUNCTIONAL EQUATIONS FOR PROBABILITY FUNCTIONS

4.1. For Lebesgue measure. Let us recall the setting for the boundary maps as in Example 1 and the functional equation (3), the second branch of the boundary maps is linear, but now we do not require the first branches of our maps to be quadratic.

We choose $0 < \hat{a}$ such that

$$\tau'_{21}(x) \leq \sigma < 1$$
 on $[0, \hat{a}], \hat{a} < a$

Note that σ depends on the choice of \hat{a} . Let us define an affine operator

$$(\mathcal{P}p)(x) = p(\tau_{21}(x)) \cdot \tau'_{21}(x) - [\tau'_{21}(x) - a \cdot \tau'_{1}(x)] .$$

The most natural space on which to consider the operator is $L^{\infty}[0, a]$, the space of bounded functions on the interval [0, a]. Note that \mathcal{P} preserves the integral on [0, 1] with respect to Lebesgue measure.

First, we will consider \mathcal{P} on the smaller space $L^{\infty}[0, \hat{a}]$.

Proposition 1. \mathcal{P} is a contraction on $L^{\infty}[0, \hat{a}]$.

Proof. We have

$$\begin{aligned} \|\mathcal{P}p - \mathcal{P}q\| &= \sup_{x \in [0,\hat{a}]} |p(\tau_{21}(x)) \cdot \tau'_{21}(x) - q(\tau_{21}(x)) \cdot \tau'_{21}(x)| \\ &\leq \sup_{x \in [0,\hat{a}]} |p(\tau_{21}(x)) - q(\tau_{21}(x))| \cdot \sup_{x \in [0,\hat{a}]} |\tau'_{21}(x)| \leq \lambda \|p - q\| . \end{aligned}$$

Proposition 2. Equation (3), which applies to selectors with $pdf \equiv 1$, considered on $[0, \hat{a}]$ has exactly one solution.

Proof. This follows by Banach's contraction principle.

Proposition 3. Using Theorem 2.2.1 in [22], the solution on $[0, \hat{a}]$ is given by

$$p(x) = \sum_{n=0}^{\infty} B(\tau_{21}^n(x)) \cdot (\tau_{21}^n)'(x) , \qquad (8)$$

where $B(x) = -[\tau'_{21}(x) - a\tau'_{1}(x)] = a\tau'_{1}(x) - \tau'_{21}(x).$

Proof. The solution p can be obtained as a limit of the functions $\mathcal{P}^n(1)$ [10]. The series (8) converges since it is dominated by a geometric series in σ .

On the other hand, we can just substitute the series (8) into (3) to get:

$$\sum_{n=0}^{\infty} B(\tau_{21}^n(x)) \cdot (\tau_{21}^n)'(x) = \sum_{n=0}^{\infty} B(\tau_{21}^n(\tau_{21}(x))) \cdot (\tau_{21}^n)'(\tau_{21}(x)) \cdot \tau_{21}'(x) + B(x) ,$$

obtaining the desired equality.

Remarks on the series (8): On every subinterval $[0, s] \subset [0, a)$ the convergence is uniform (for $\hat{a} < s$ as well). The series is divergent at a (unless B(a) = 0) so the value of p(a) should be obtained using $p(\tau_{21}(a))$ and equation (3). Let us define

$$f_k(x) = \sum_{n=0}^k B(\tau_{21}^n(x)) \cdot (\tau_{21}^n)'(x) , \quad k = 0, 1, 2, \dots$$

Note that $\int_0^a f_k(x)dx = 0$ for all $k \ge 0$, while for the limit p we usually have $\int_0^a p(x)dx > 0$. This, in general, means that the functions f_k are not uniformly integrable (i.e., for arbitrary constant M > 0 and arbitrary $\varepsilon > 0$ we can find $k \ge 0$ such that $\int_{[a-\varepsilon,a]} |f_k(x)|dx > M$) and the series in (8) does not converge in $L^1[0,a]$.

Proposition 4. Once the solution p is known on $[0, \hat{a}]$ it is uniquely extended to [0, a). The solution is still described by the formula (8). The value of p(a) is obtained using $p(\tau_{21}(a))$ and equation (3).

Proof. First note that the function τ_{21}^{-1} has two fixed points, 0 and a. Since $\tau_{21}(x) < x$ on (0, a), the sequence $\{\tau_{21}^{-n}(\hat{a})\}$ is strictly increasing and converges to the fixed point a as $n \to \infty$. We extend p using equation (3): from $[\tau_{21}(\hat{a}), \hat{a})$ to $[\hat{a}, \tau_{21}^{-1}(\hat{a}))$, then to $[\tau_{21}^{-2}(\hat{a}), \tau_{21}^{-2}(\hat{a}))$, then to $[\tau_{21}^{-2}(\hat{a}), \tau_{21}^{-3}(\hat{a}))$, etc.

The solution is uniquely described by the formula (8).

Remark 1. Note that p is uniquely determined and independent of the choice of \hat{a} .

Example 4 shows a successful application of Propositions 1-4 to find the probability function defining a random map with a required pdf. Example 5 shows that the method of Propositions 1-4 sometimes fails. The function produced is not a probability function.

Example 4. We set a = 1/2. We consider $\tau_1(x) = 2x^2 + x$, $\tau_2(x) = -2x^2 + 3x$, $\tau_{21}(x) = \tau_2^{-1}(\tau_1(x)) = 3/4 - (1/4)\sqrt{9 - 16x^2 - 8x}$. Using Maple 13 we were able to guess that the solution is p(x) = x + 1/4. We illustrate the statements with a number of pictures depicted in Figs.6-9.

Example 5. In this example the solution p is not between 0 and 1. We set a = 1/5 and consider

$$\tau_1(x) = \begin{cases} (4/3)x^2 + 1.365128205x &, \text{ for } 0 \le x < 0.13 ; \\ 138.8110936x^2 - 34.37908948x + 2.323374150 &, \text{ for } 0.13 \le x \le a , \end{cases}$$

and $\tau_2(x) = 1 - \tau_1(1/5 - x)$, $\tau_{21}(x) = \tau_2^{-1}(\tau_1(x))$. We illustrate the statements with a number of pictures depicted in Figs.10-13.



FIGURE 8. Derivative τ'_{21} for Example 4.

FIGURE 9. First five iterations of approximating the invariant solution p on [0, 0.5]. Cut off at 3 for Example 4.

4.2. General density functions. We now consider the case when the selector τ has a general pdf. Let us consider two maps $\tau_1 : [0, a] \to [0, 1]$ increasing with $\tau_1(0) = 0$, $\tau_1(a) = 1$ satisfying $\tau_1(x) \le x/a$, and $\tau_{21} : [0, a] \to [0, a]$ increasing with $\tau_{21}(0) = 0$, $\tau_{21}(a) = a$ satisfying $\tau_{21}(x) < x$ on open interval (0, a). Both τ_i are extended onto [0, 1] by defining them on (a, 1] as a monotonic map $\tau : (a, 1] \to [0, 1]$ which is onto and expanding (i.e. $|\tau'| > 1$).

Let f be an invariant density of a selector whose graph is between the graphs of τ_1 and τ_2 . We are looking for a probability p(x) such that the random map $R = \{\tau_1, \tau_2; p, 1 - p\}$ preserves the density f. The corresponding Frobenius-Perron equation is

$$f(x) = \frac{p(\phi_1(x))f(\phi_1(x))}{\tau_1'(\phi_1(x))} + \frac{(1 - p(\psi_1(x)))f(\psi_1(x))}{\tau_2'(\psi_1(x))} + \frac{f(\tau^{-1}(x))}{|\tau'(\tau^{-1}(x))|},$$
(9)



FIGURE 12. Derivative τ'_{21} for Example 5.

FIGURE 13. First three iterations of approximating the invariant solution p on [0, 0.2], cutting off between -7 and 5 for Example 5.

where $\phi_1 = (\tau_1|_{[0,a]})^{-1}$, $\psi_1 = (\tau_2|_{[0,a]})^{-1}$, $x \in [0,1]$. Introducing $g = p \cdot f$, substituting $x = \tau_1(y)$, $y \in [0,a]$ and using the equality

$$\tau_{21}'(y) = \frac{\tau_1'(y)}{\tau_2'(\tau_{21}(y))} ,$$

we reduce eq.(9) to

$$g(y) = g(\tau_{21}(y)) \cdot \tau'_{21}(y) + \left[\left(\left(f(\tau_1(y)) - \frac{f(\tau^{-1}(\tau_1(y)))}{|\tau'(\tau^{-1}(\tau_1(y)))|} \right) \cdot \tau'_1(y) - \tau'_{21}(y) \right] .$$
(10)

We want to solve this equation hoping that p will satisfy 0 .

We choose $\sigma>0$ such that

$$\tau'_{21}(x) \le \sigma < 1$$
 , on interval $[0, \widehat{a}]$, $\widehat{a} < a$.

Let us define the affine operator

$$(\mathcal{P}g)(x) = g(\tau_{21}(y)) \cdot \tau'_{21}(y) + \left[\left((f(\tau_1(y)) - \frac{f(\tau^{-1}(\tau_1(y)))}{|\tau'(\tau^{-1}(\tau_1(y)))|} \right) \cdot \tau'_1(y) - \tau'_{21}(y) \right]$$

The most natural space to consider it on is the space $L^{\infty}[0, a]$ of bounded functions on the interval [0, a].

First, we will consider \mathcal{P} on the smaller space $L^{\infty}[0, \hat{a}]$. The proofs of Proposition 5 and Corollary 2 are repetition of those for Propositions 1-4.

Proposition 5. \mathcal{P} is a contraction on $L^{\infty}[0, \widehat{a}]$. Its unique fixed point is given by

$$g(x) = \sum_{n=0}^{\infty} B(\tau_{21}^n(x)) \cdot (\tau_{21}^n)'(x) , \qquad (11)$$

where $B(y) = \left(f(\tau_1(y)) - \frac{f(\tau^{-1}(\tau_1(y)))}{|\tau'(\tau^{-1}(\tau_1(y)))|} \right) \cdot \tau'_1(y) - \tau'_{21}(y).$

Corollary 2. There exists a unique function p(x) such that eq.(9) holds. On $[0, \hat{a})$, p(x) is given by

$$p(x) = \frac{1}{f(x)} \sum_{n=0}^{\infty} B(\tau_{21}^n(x)) \cdot (\tau_{21}^n)'(x)$$
(12)

and it can be extended to [0, a] as described in Proposition 4.

5. Singular + Singular = Lebesgue

In this Section we use the results of Example 1 to construct maps τ_1 and τ_2 that have no acims and a probability function p(x) such that the random map $R = \{\tau_1, \tau_2; p, 1-p\}$ preserves the Lebesgue measure.

Let us consider maps $\tau_i : [0, 1/2] \to [0, 1], i = 1, 2$, increasing, with $\tau_i(0) = 0, \tau_i(1/2) = 1$ satisfying $\tau_1(x) \le 2x, \tau_2(x) \ge 2x$ and the map $\tau_{21} = \tau_2^{-1} \circ \tau_1 : [0, 1/2] \to [0, 1/2]$ increasing with $\tau_{21}(0) = 0, \tau_{21}(1/2) = 1/2$ satisfying $\tau_{21}(x) < x$ on open interval (0, 1/2).

We consider the special case of quadratic τ_1 and τ_2 . Let

$$\tau_1(x) = a_1 x^2 + b_1 x$$
, $\tau_2(x) = a_2 x^2 + b_2 x$,

where $a_i = 4(1 - b_i/2), i = 1, 2, 0 < b_1 < 2 < b_2 < 4$. We have

$$\tau_{21}(x) = \frac{1}{2(1-b_2/2)} \cdot \left(-b_2/4 + \sqrt{b_2^2/16 + 4(1-b_2/2)[(1-b_1/2)x^2 + b_1x/4]} \right)$$

We extend the maps τ_1 and τ_2 to the interval [0, 1] as follows

$$\tau_i(x) = \begin{cases} \tau_i(x), & \text{for } 0 \le x \le 1/2 ; \\ \tau_i(x-1/2), & \text{for } 1/2 < x \le 1 , \end{cases}$$

i = 1, 2. We want to find a probability p(x) such that the random map $R = \{\tau_1, \tau_2; p, 1-p\}$ preserves Lebesgue measure. The corresponding Frobenius-Perron equation is

$$1 = \frac{p(\phi_1(x))}{\tau_1'(\phi_1(x))} + \frac{1 - p(\psi_1(x))}{\tau_2'(\psi_1(x))} + \frac{p(\phi_1(x) + 1/2)}{\tau_1'(\phi_1(x))} + \frac{1 - p(\psi_1(x) + 1/2)}{\tau_2'(\psi_1(x))} ,$$
(13)

where $\phi_1 = (\tau_1|_{[0,1/2]})^{-1}$, $\psi_1 = (\tau_2|_{[0,1/2]})^{-1}$, $x \in [0,1]$.

We will look for a solution p such that p(x + 1/2) = p(x) for $x \in [0, 1/2]$. Then the equation (13) reduces to

$$\frac{1}{2} = \frac{p(\phi_1(x))}{\tau_1'(\phi_1(x))} + \frac{1 - p(\psi_1(x))}{\tau_2'(\psi_1(x))} , \qquad (14)$$

 $x \in [0,1]$. Substituting $x = \tau_1(y), y \in [0,1/2]$ and using the equality

$$au_{21}'(y) = rac{ au_1'(y)}{ au_2'(au_{21}(y))} ,$$



for $b_1 = 0.5, b_2 = 3.5$.



we reduce it further to the known equation

$$p(y) = p(\tau_{21}(y)) \cdot \tau'_{21}(y) - [\tau'_{21}(y) - (1/2) \cdot \tau'_{1}(y)] \quad . \tag{15}$$

Using Maple 13 we previously found that the linear function p(x) = A + Bx satisfies (15) if we put

$$A = \frac{b_1(1 - b_2/2)}{b_1 - b_2} \quad , \quad B = \frac{8(1 - b_1/2)(1 - b_2/2)}{(b_1 - b_2)}$$

Note, that both coefficients are positive as $b_2 > b_1$, $b_1/2 < 1$ and $b_2/2 > 1$.

For $b_1 = 0.5$, $b_2 = 3.5$ neither map τ_1 nor τ_2 has an acim. Almost all trajectories converge to 0 or 1 correspondingly. However, the random map $R = \{\tau_1, \tau_2; p, 1-p\}$, where

$$p(x) = \begin{cases} (3/2)x + 1/8, & \text{for } 0 \le x \le 1/2 ; \\ (3/2)(x - 1/2) + 1/8, & \text{for } 1/2 < x \le 1 , \end{cases}$$

does preserve Lebesgue measure.

6. Invariant density of any random map is also an invariant density of a selector

Let τ_1 and τ_2 , lower and upper maps, respectively, be in \mathcal{T} and possess acims with densities f_1 and f_2 . It is shown in [15] that \mathcal{D} , the set of densities invariant under the position dependent random maps $R = \{\tau_1, \tau_2; p, 1-p\}$, is a convex set. Since the underlying space I = [0, 1] is compact, the space of probability measures on I is weakly compact [27, Theorem 6.4]. From Theorem 6 of [15] it follows that \mathcal{D} is weakly closed and hence weakly compact in L^1 .

Let BV(I) be the space of functions of bounded variation on I. It is proved in [6] that if τ_1 and τ_2 satisfy the following two conditions

$$\begin{aligned} (A) \ &\sum_{k=1}^2 g_k(x) < \alpha < 1, \ \text{for some } \alpha, \\ (B) \ &g_k \in BV(I), k = 1, 2, \end{aligned}$$

where $g_k(x) = \frac{p_k(x)}{|\tau'_k(x)|}$, k = 1, 2, then the random map $R = \{\tau_1, \tau_2; p, 1-p\}$ has an invariant density of bounded variation. Since we are considering piecewise expanding maps,

$$\sum_{k=1}^{2} g_k(x) < \sum_{k=1}^{2} p_k(x) = 1$$

This implies that condition (A) is always satisfied. If we collect all possible probability functions p such that the above condition (B) is satisfied uniformly, then the corresponding set of invariant densities, \mathcal{D}_{AB} , is a bounded set in BV(I). Thus, the L^1 compactness of \mathcal{D}_{AB} is established.

Consider the convex hull of f_1 and f_2 , denoted by $\mathcal{D}_{Co} = \{\alpha f_1 + (1 - \alpha)f_2 | 0 \le \alpha \le 1\}$. It follows from Corollary 1 that we can choose p such that $R = \{\tau_1, \tau_2; p, 1 - p\}$ preserves any $f \in \mathcal{D}_{Co}$. \mathcal{D}_{Co} is closed and compact, has two extreme points, f_1 and f_2 , and is a "line segment" connecting f_1 and f_2 .

We define

 $\mathcal{D}_S = \{f | f \text{ is an invariant pdf of some selector } \boldsymbol{\tau} \in \mathcal{T} \}.$ $\mathcal{D}_{RM} = \{f | f \text{ is an invariant pdf of some random map } R = \{\tau_1, \tau_2; p, 1-p\} \}.$ From Examples 3 and 5, we know

 $\mathcal{D}_S \not\subseteq \mathcal{D}_{RM}.$

By Theorem 1, we have

 $\mathcal{D}_{Co} \subset \mathcal{D}_S.$

From Corollary 1, it follows that

$$\mathcal{D}_{Co} \subset \mathcal{D}_{RM}$$

We will now prove that $\mathcal{D}_{RM} \subset \mathcal{D}_S$, assuming that we consider only random maps whose pdf's are positive a.e. on [0, 1].

Theorem 3. Let us consider two maps $\tau_k : [0,1] \to [0,1]$, k = 1,2, having the same monotonicity on the intervals of monotonicity partition, piecewise expanding (we assume slope > 2, or harmonic average of slopes condition [18]). Then $\mathcal{D}_{RM} \subset \mathcal{D}_S$ if we consider only random maps whose pdf's are positive a.e. on [0,1].

Proof. We will use the notations introduced at the beginning of Section 3. Recall the definition of extended inverse function in the Section 3. Assuming that f is an invariant density of the random map $R = \{\tau_1, \tau_2; p, 1-p\}$, the Frobenius-Perron equation for R is

$$f(x) = \sum_{j=1}^{m} \frac{p(\tau_{1,j}^{-1}(x))f(\tau_{1,j}^{-1}(x))}{|\tau_{1}'(\tau_{1,j}^{-1}(x))|} \chi_{\tau_{1,j}(I_{j})}(x) + \frac{(1 - p(\tau_{2,j}^{-1}(x)))f(\tau_{2,j}^{-1}(x))}{|\tau_{2}'(\tau_{2,j}^{-1}(x))|} \chi_{\tau_{2,j}(I_{j})}(x).$$
(16)

Let $\mathbf{F}(x) = \int_0^x f(t) dt$, $x \in I$. Now construct a map $\boldsymbol{\tau}$, piecewise as follows:

$$\overline{\boldsymbol{\tau}_{j}^{-1}}(x) = \mathbf{F}^{-1}\left(\int_{0}^{\overline{\boldsymbol{\tau}_{1,j}^{-1}}(x)} p(t)f(t)\,dt + \int_{0}^{\overline{\boldsymbol{\tau}_{2,j}^{-1}}(x)} (1-p(t))f(t)\,dt\right).$$
(17)

Note that τ defined above has the same monotonicity as the boundary maps after the vertical segments are removed. Equation (17) is equivalent to

$$\mathbf{F}\left(\overline{\boldsymbol{\tau}_{j}^{-1}}(x)\right) = \int_{0}^{\overline{\boldsymbol{\tau}_{1,j}^{-1}}(x)} p(t)f(t) \, dt + \int_{0}^{\overline{\boldsymbol{\tau}_{2,j}^{-1}}(x)} (1-p(t))f(t) \, dt \, .$$

Differentiating both sides of the above equation with respect to x, we have

$$\frac{f(\tau_{j}^{-1}(x))}{\mid \tau'(\tau_{j}^{-1}(x))\mid}\chi_{\tau_{j}(I_{j})}(x) = \frac{p(\tau_{1,j}^{-1}(x))f(\tau_{1,j}^{-1}(x))}{\mid \tau_{1}'(\tau_{1,j}^{-1}(x))\mid}\chi_{\tau_{1,j}(I_{j})}(x) + \frac{(1-p(\tau_{2,j}^{-1}(x)))f(\tau_{2,j}^{-1}(x))}{\mid \tau_{2}'(\tau_{2,j}^{-1}(x))\mid}\chi_{\tau_{2,j}(I_{j})}(x).$$
(18)

Now, from the definition of the Frobenius-Perron operator of τ , P_{τ} , it follows

$$(P_{\tau}f)(x) = \sum_{j=1}^{m} \frac{f(\tau_{j}^{-1}(x))}{|\tau'(\tau_{j}^{-1}(x))|} \chi_{\tau_{j}(I_{j})}(x)$$

$$= \sum_{j=1}^{m} \frac{p(\tau_{1,j}^{-1}(x))f(\tau_{1,j}^{-1}(x))}{|\tau'_{1}(\tau_{1,j}^{-1}(x))|} \chi_{\tau_{1,j}(I_{j})}(x) + \frac{(1 - p(\tau_{2,j}^{-1}(x)))f(\tau_{2,j}^{-1}(x))}{|\tau'_{2}(\tau_{2,j}^{-1}(x))|} \chi_{\tau_{2,j}(I_{j})}(x)$$

$$= f(x),$$

$$(19)$$

which implies that τ preserves f. We have $\overline{\tau_{2,j}^{-1}}(x) \leq \overline{\tau_{1,j}^{-1}}(x)$, so we also have

$$\begin{aligned} \mathbf{F}\left(\overline{\boldsymbol{\tau}_{j}^{-1}}(x)\right) &\leq \int_{0}^{\overline{\boldsymbol{\tau}_{1,j}^{-1}}(x)} p(t)f(t)\,dt + \int_{0}^{\overline{\boldsymbol{\tau}_{1,j}^{-1}}(x)} (1-p(t))f(t)\,dt \\ &= \int_{0}^{\overline{\boldsymbol{\tau}_{1,j}^{-1}}(x)} f(t)\,dt = \mathbf{F}\left(\overline{\boldsymbol{\tau}_{1,j}^{-1}}(x)\right), \end{aligned}$$

and

$$\begin{split} \mathbf{F}\left(\overline{\boldsymbol{\tau}_{j}^{-1}}(x)\right) &\geq \int_{0}^{\overline{\boldsymbol{\tau}_{2,j}^{-1}}(x)} p(t)f(t)\,dt + \int_{0}^{\overline{\boldsymbol{\tau}_{2,j}^{-1}}(x)} (1-p(t))f(t)\,dt \\ &= \int_{0}^{\overline{\boldsymbol{\tau}_{2,j}^{-1}}(x)} f(t)\,dt = \mathbf{F}\left(\overline{\boldsymbol{\tau}_{2,j}^{-1}}(x)\right), \end{split}$$

therefore,

$$\mathbf{F}\left(\overline{\tau_{2,j}^{-1}}(x)\right) \leq \mathbf{F}\left(\overline{\tau_{j}^{-1}}(x)\right) \leq \mathbf{F}\left(\overline{\tau_{1,j}^{-1}}(x)\right).$$

Then, it follows from the definition of \mathbf{F} that

$$\tau_{1,j}(x) \le \boldsymbol{\tau}_j(x) \le \tau_{2,j}(x).$$

This implies that τ is a selector. The proof is completed.

7. CONTINUOUS DEPENDENCE OF INVARIANT DENSITIES ON POSITION DEPENDENT RANDOM MAPS

The following continuity theorem will be used in the next section.

Theorem 4. Let us consider two maps $\tau_k : [0,1] \to [0,1], k = 1,2$, both piecewise expanding (we assume slope > 2, or harmonic average of slopes condition [18]). Let us consider a sequence of piecewise semi-Markov approximations $\tau_k^{(n)}$ on uniform partitions \mathcal{P}_n , $\tau_k^{(n)} \to \tau_k$ in L^1 as $n \to \infty$. Let $R^{(n)} = \{\tau_1^{(n)}, \tau_2^{(n)}; p_1^{(n)}, p_2^{(n)}\}$ be a position dependent random map preserving density $f_{(n)}$. If $p_1^{(n)} \to p_1$ pointwise almost everywhere (actually even weak L^{∞} convergence suffices), then $f_{(n)} \to f$, where f is invariant for $R = \{\tau_1, \tau_2; p_1, p_2\}, p_1 + p_2 = 1.$

Proof. It follows from the slope assumption that the Frobenius-Perron operators $P_{R^{(n)}}$ satisfy the Lasota-Yorke inequality [10,24] with uniform constant. And thus, $\{f_{(n)}\}\$ is a precompact set in L^1 . We assume that the limit of $\{f_{(n)}\}$ is f.

Let us introduce one more sequence of random maps: $\widehat{R}^{(n)} = \left\{\tau_1, \tau_2; p_1^{(n)}, p_2^{(n)}\right\}$, their invariant densities are denoted by $\widehat{f}_{(n)}, n \in \mathbb{N}$.

$$\| P_R f - f \|_{L^1} \leq \| P_R f - P_{\widehat{R}^{(n)}} f \|_{L^1} + \| P_{\widehat{R}^{(n)}} f - P_{R^{(,n)}} f_{(n)} \|_{L^1} + \| P_{\widehat{R}^{(n)}} f_{(n)} - P_{R^{(n)}} f_{(n)} \|_{L^1} + \| P_{R^{(n)}} f_{(n)} - f \|_{L^1} = \| \sum_{k=1}^2 P_{\tau_k} p_k f - \sum_{k=1}^2 P_{\tau_k} p_k^{(n)} f \|_{L^1} + \| P_{\widehat{R}^{(n)}} (f - f_{(n)}) \| + \| P_{\widehat{R}^{(n)}} f_{(n)} - P_{R^{(n)}} f_{(n)} \|_{L^1} + \| f_{(n)} - f \|_{L^1} \leq \sum_{k=1}^2 \| (p_k - p_k^{(n)}) f \|_{L^1} + \| f - f_{(n)} \|_{L^1} + \| P_{\widehat{R}^{(n)}} f_{(n)} - P_{R^{(n)}} f_{(n)} \|_{L^1} + \| f_{(n)} - f \|_{L^1} .$$

The first summand is sufficiently small when n large enough since $p_k^{(n)} \to p_k$. The second and fourth terms approach 0 by the definition. For the third term, note that $R^{(n)}$ and $\hat{R}^{(n)}$ have the same probability functions, we also have $\tau_k^{(n)} \to \tau_k$ as $n \to \infty$, it follows from the stability of $\hat{R}^{(n)}$ that this term also approaches 0 as $n \to \infty$. This completes the proof.

Remark 2. We need the assumption that slope > 2 or harmonic average of slopes condition for τ_k . For example, $\tau_1^{(n)} = W_0$, $\tau_2^{(n)} = W_{a_n}$, where W_0 is the original W-shaped map and W_{a_n} is a sequence of W-shaped maps of W_0 being perturbed in a neighborhood of 1/2, call it $U_{(1/2)}$ (The detailed construction of W_{a_n} can be found in [25]). We put $p_2^{(n)} = 1$ in the set $U_{(1/2)}$. Then as $R^{(n)} = \{W_0, W_{a_n}; p_1^{(n)}, p_2^{(n)}\} \rightarrow R = \{W_0, W_0; p_1, p_2\}, f_{(n)}$ shall converge to a measure with singular component. However, R preserves some acim.

8. Optimization and extreme points

A well known formulation of the Krein-Milman theorem states that every linear continuous functional on a locally convex Hausdorff linear space E attains its minimum on a compact subset X of Eat some extreme point of X [9]. This result does not require compactness of X and has been extended to lower semi-continuous concave functionals in [7].

In general we do not know the extreme points of \mathcal{D}_{RM} . However, in the special case where τ_1 and τ_2 are semi-Markov piecewise linear, it has been shown in [16] that the extreme points of \mathcal{D}_{RM} come from the random maps where, for each $x \in I$, $p(x) \in \{0, 1\}$, that is, deterministic maps, taking their values on either the lower or upper boundaries. We call such a random map a bang-bang map. In this setting, we consider a continuous linear functional on \mathcal{D}_{RM} which is to be optimized. By the Krein-Milman theorem the random map that optimizes the functional over the admissible probability density functions is a bang-bang map.

A typical continuous linear functional that appears in optimization problems is of the form:

$$F(f) = \int_0^1 g(x)f(x)dx,$$

where g is a fixed bounded function on I and f is a density function in \mathcal{D}_{RM} .

Question: Can we generalize the above bang-bang result to more general maps?

From now on, let us consider the real space L_I^1 of Lebesgue integrable function on I. It is a locally convex, linear Hausdorff space. For a sequence of closed sets of L_I^1 , $\{S_n\}$, we write $\limsup S_n$ to refer to the limit superior in the sense of Kuratowski convergence [23].

Definition 1. The point **s** belongs to $\limsup S_n$, if every neighbourhood of **s** intersects an infinite number of the S_n .

We will also need the following theorem from [20, Theorem 2], which we have modified for our use.

 L^1

Theorem 5. Let $\{S_n\}$ be a sequence of compact convex sets whose union is contained in a compact convex subset of L_I^1 . Let \mathcal{X}_n be the set of extreme points of \mathcal{S}_n for each n and let $\mathcal{S} = \limsup \mathcal{S}_n$ and $\mathcal{X} = \limsup \mathcal{X}_n$. Then $\mathcal{S} \subset Co(\mathcal{X})$, where $Co(\mathcal{X})$ is the convex hull of \mathcal{X} . If, in addition, \mathcal{S} is convex, then $\mathcal{S} = Co(\mathcal{X})$, that is, \mathcal{X} contains all the extreme points of \mathcal{S} .

Our boundary maps are assumed to be piecewise expanding, it follows from Pelikan's condition [28] that the random map $R = \{\tau_1, \tau_2; p, 1-p\}$ always has an invariant density function in L_I^1 for any measurable probability function p. We introduce two sequences of semi-Markov maps, $\{\tau_1^N\}$ and $\{\tau_2^N\}$, which approximate τ_1 and τ_2 respectively on uniform partitions of I into N equal subintervals $\{I_i^u\}_{i=1}^N$. Moreover, we approximate the probability function p by a sequence of step functions on these subintervals, denoted by $\{p^N\}$, i.e. $p^N = \sum_{i=1}^N p_i^N \chi_{I_i^u}$. We can do this in such way that $p^N \to p$ pointwise almost everywhere.

We define sets of attainable densities and consider them as subsets of L^1 :

 $\mathcal{A} = \{f : \text{ there exits probability } p \text{ such that } f \text{ is an invariant density of } R = \{\tau_1, \tau_2; p, 1-p\}\},\$

 $\mathcal{A}_N = \left\{ f_N : \text{ there exits probability } p^N \text{ which is piecewise constant on the uniform partition} \right.$

 $\{I_i^u\}_{i=1}^N \text{ such that } f_N \text{ is an invariant density of } R^N = \{\tau_1^N, \tau_2^N; p^N, 1-p^N\} \},$

 $\mathcal{E}_N = \{f : f \text{ is an extreme point in } \mathcal{A}_N\},\$

$$\mathcal{E} = \limsup \mathcal{E}_N,$$

where $N \in \mathbb{N}$. Note that each \mathcal{A}_N is defined for fixed τ_1^N and τ_2^N , and p^N can be viewed as a vector which varies.

Remark 3. (I) \mathcal{A}_N is convex [16]. Since τ_1^N and τ_2^N have large enough slope, \mathcal{A}_N is set of functions of bounded variations, and therefore is precompact. Furthermore, it follows from Lemma 1 and Theorem 6 in [15] that, \mathcal{A}_N is closed. Therefore \mathcal{A}_N is compact.

(II) Each \mathcal{A}_N is contained in a set of functions of bounded variations and it contains all density functions which are the approximations of the invariant density functions of the random map $R = \{\tau_1, \tau_2; p, 1-p\}$ for some p. Thus $\{\mathcal{A}_N\}$ is contained in some closed set of functions of bounded variations, which is compact.

(III) Now, consider a density function f which is invariant for the random map $R = \{\tau_1, \tau_2; p, 1-p\}$ for some p. It follows from Theorem 4 that we can choose a sequence of maps $R^N = \{\tau_1^N, \tau_2^N; p^N, 1-p^N\}$ such that

$$f^N \in \mathcal{A}_N, \ \tau_1^N \to \tau_1, \ \tau_2^N \to \tau_2, \ p^N \to p,$$

where f^N is the invariant density of \mathbb{R}^N , and f^N converges to the invariant density, f, of \mathbb{R} . The convergence implies every neighbourhood of f meets the sets of the sequence $\{\mathcal{A}_N\}$ with arbitrarily large index N. Thus,

$$\limsup \mathcal{A}_N = \mathcal{A}.$$

 $(IV) \mathcal{A} \text{ is convex } [15, 16].$

It follows from Theorem 5 and (I-IV) in the Remark 3 that:

Theorem 6. $Co(\mathcal{E}) = \mathcal{A}$, where $Co(\mathcal{E})$ is the convex hull of \mathcal{E} .

This means \mathcal{E} contains all the extreme points of \mathcal{A} .

Corollary 3. Every invariant density which is an extreme point in \mathcal{A} is a limit of densities f^N which are invariant for some random map \mathbb{R}^N with probability $p^N(x) \in \{0,1\}$.

Corollary 3 suggests the following algorithm to optimize a continuous functional F on the set of densities \mathcal{D}_{RM} . For large enough N, construct invariant densities f_k^N for all "bang-bang" probabilities p_k^N . Find f_k^N which optimizes F on \mathcal{D}_{RM}^N . Then, $\{F(f_k^N)\}$ converges to the optimal value of F on \mathcal{D}_{RM} .

9. Summary of examples

(i) Example 1 presents a positive general solution for the case when the selector is the triangle map and the boundary maps have two branches, the first quadratic, the second linear.

(ii) Example 2 presents a positive solution for the specific case of a selector and semi-Markov boundary maps with five branches.

(iii) Example 3 shows that with the same boundary maps as in Example 2 and another selector (which has an invariant pdf), it is impossible.

(iv) Example 4 demonstrates a successful application of Propositions 1-4 to find the probability function defining a random map with the required pdf.

(v) Example 5 shows that the method of Propositions 1-4 sometimes fails. The function produced is not a probability function.

(vi) In Section 5 we produce boundary maps that have no acim, but a random map based on them preserves Lebesgue measure invariant for a selector.

10. Concluding Remarks

The main objective of this paper is to study conditions for a selector of a multivalued function with graph G to have statistical dynamics that can be represented by the dynamics of a position dependent random map based solely on the boundaries of G. Theorem 1 and Corollary 1 state that any convex combination of pdf's of boundary maps can be realized both as the pdf of a selector and as the pdf of a random map. We develop results that work in general but demonstrate with examples that some of our methods depend very sensitively on the selector and on the boundary maps.

We also study the converse problem: when can a pdf of a random map based on the boundary maps be realized as a pdf of some selector? Theorem 3 proves that this always holds in the case of piecewise expanding boundary maps which are piecewise monotonic having the same monotonicity on the partition intervals. Finally we study the extreme points of the set of pdf's of all random maps based on the boundary maps and attempt to characterize them.

In the future we plan to generalize Theorem 3 omitting, if possible, the assumption of common monotonicity on partition intervals. A major project will be the generalization of our results to higher dimensions.

Acknowledgement. We are extremely grateful for the very helpful comments of the anonymous reviewers. Their suggestions and critiques have greatly improved the paper.

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